On the Structure of Tenancy Contracts: Theory and Evidence from 19th Century Rural Sicily^{*}

Oriana Bandiera London School of Economics and Political Science

The Suntory Centre

Suntory and Toyota International Centres for

Economics and Related Disciplines

London School of Economics and Political Science

Houghton Street London WC2A 2AE

Tel: (020) 7955 6674

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Abstract

This paper analyses the empirical determinants of contract length, a key and yet neglected dimension of contractual structure. I use data on tenancy agreements signed between 1870 and 1880 in the district of Siracusa, Italy to estimate the choice over length and compensation schemes jointly.

The findings indicate that the choice of contract length is driven by the need to provide incentives for non-observable investment, taking into account transaction costs and imperfections in the credit markets that make incentive provision costly. The results also illustrate that since both length and the compensation scheme are used to provide incentives within the same contract, joint analysis is important for a correct interpretation of the evidence.

Keywords: contract duration, incentives, tenancy agreements.

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<u>Contact address:</u> Dr Oriana Bandiera, STICERD, London School of Economics and Political Science, Houghton Street, London WC2A 2AE, UK. Email: o.bandiera@lse.ac.uk

1 Introduction.

Contracts regulate most economic transactions and, not surprisingly, contract theory is a cornerstone of economics. In contrast to the large body of theory, evidence on the determinants of contractual structure is limited and typically focuses on the compensation scheme only.¹

The aim of this paper is to present evidence on the determinants of the compensation scheme joint with contract *duration*, a key and yet neglected dimension of contractual structure.

Analyzing contract duration is of interest for two reasons. First, in a principal-agent framework where the agent works with long-lived assets, contract duration determines the agent's stake in future production and hence the incentive to undertake non observable investments. Evidence on the determinants of duration can thus shed light on the extent to which incentives for non observable investments are provided in practice.

Second, since both contract duration and the compensation scheme are used to provide incentives within the same contract, studying them jointly is key to provide an accurate picture of the determinants of contract form. Focusing on one dimension only can, in contrast, mislead the interpretation of the evidence.

This paper uses new data on land tenancy agreements signed between 1870 and 1880 in the district of Siracusa, Italy. The data set, which I built from the original contract documents, contains information on both duration and the compensation scheme. As usual in agriculture, the latter is either sharecropping, whereby the tenant and the landlord share the output, or fixed rent, whereby the tenant retains the output and pays a fixed rent to the landlord.

The compensation scheme determines incentives for non observable production effort. While under both fixed rent and sharecropping contracts the tenant's pay depends on the realization of output, fixed rent contracts give the tenant a higher stake in production and hence provide stronger effort incentives.

The length of the contract, on the other hand, determines incentives for non

¹Chiappori and Salanie [2003] survey the recent empirical literature on the determinants of contractual structure.

observable *investment* effort. Contract duration determines whether the tenant has a stake in future output and therefore the incentive to undertake non observable investment that increases output in future periods.

The data contains information on crops', tenants' and landlords' characteristics that relate directly to the benefits and costs of providing incentives. Crops in the sample differ in their riskiness and their sensitivity to investment effort. Tenants' wealth affects their risk aversion and the extent to which they are subject to limited liability. Landlords' characteristics, namely their gender, their social class and whether they live close to the rented plot affect transaction costs.

The empirical findings support the idea that contracts are designed to provide incentives, taking into account transaction costs and imperfections in the credit and insurance markets. I find that high powered incentives for both effort and investment are provided when the cost of doing so is low, that is when the tenant is rich. If incentive provision is costly, as it is for poor tenants, high powered investment incentives via long term contracts are only offered when the benefit is high enough, namely for trees that are more sensitive to investment effort.

Landlords who face higher monitoring and renegotiation costs, namely female and aristocratic landlords who are much less likely to be daily in the fields, choose the contract combination that minimizes these costs: long term with fixed rent. Also, these landlords are less likely to ever want to resume direct cultivation and hence place less value on the flexibility given by short term contracts.

Overall, the findings are consistent with the idea that contract duration is key to provide incentives for non observable investment. Whether investment incentives are provided, however, depends on the crop's, the tenant's and the landlord's characteristics. The results also illustrate that analyzing one contract dimension by itself can be misleading if the effect of the exogenous variables on one dimension depends on the choice in other dimensions.

This paper contributes to the empirical contract literature by offering the first joint analysis of contract type and duration.² Moreover, in contrast to the survey

²Evidence on the determinants of contract duration *per se* is quite limited. Exceptions are Joskow [1987], who shows that contracts between coal suppliers and electric utilities are significantly longer

data generally used in the literature, original documents contain information on both contracting parties.³ Taken together, these features of the data offer a more accurate picture of the determinants of contract form.

The remainder of the paper is organized as follows. Section 2 draws ideas on the determinants of contractual structure from the existing theoretical literature. Section 3 describes the data set and the methodology. Section 4 presents the main findings. Section 5 reports extensions and discusses econometric concerns. Section 6 concludes.

2 The Determinants of Contract Structure: Ideas from Theory.

Theories of contractual structure typically analyze contractual choice in a principal-agent framework where the principal chooses the terms of the contract to maximize her payoff for given characteristics of the agent and production function.⁴ In the context under study, the principal hires the agent to cultivate her land and chooses two dimensions of contractual structure- the duration of the contract and the agent's compensation scheme.

Within the first dimension, duration, the principal chooses between short term and long term contracts. Short term contracts are one period long, where one period is defined as the length of time within which the agent performs his tasks, the outcome is realized and the agent receives a payment. In the agricultural context studied here, one period typically corresponds to one calendar year. Long term contracts are

when relationship-specific investment is important, and Crocker and Masten [1988] who show that natural gas contracts are shorter when flexibility becomes exogenously more relevant. Brickley et al [2003] show that length of franchise agreements increases with the importance of non-contractible investments and decreases when the need for flexibility increases. In line with these studies I find that investment incentives and flexibility are significant determinants of contract length.

The literature on contract type is more extensive, both in general and for the specific case of land tenancy. The findings in this paper are generally in line with existing evidence that sharecropping contracts are more likely to be offered to poor tenants and that crop characteristics are a significant determinant of contract type. See Ackerberg and Botticini [2000, 2002], Allen and Lueck [1996], Dubois [2002], Laffont and Matoussi [1995].

³Most empirical studies on tenancy agreements only have information on the tenant (e.g. Laffont and Matoussi 1995) and others only on the landlord (e.g. Dubois 2002). The data set used by Ackerberg and Botticini [2000, 2002] is an exception in this regard.

⁴The principal is thus assumed to have all the bargaining power and matching between principal and agents is assumed to be random.

agreements that last more than one period.

Within the second dimension, type, the principal chooses between sharecropping and fixed rent contracts. The key difference between the two is that under sharecropping the principal and the agent each take half of the output, while under fixed rent the agent retains the whole output and pays the principal a fixed amount at the end of each period.⁵

The choice of contractual structure is driven by three sets of considerations or characteristics of the environment under scrutiny. First are the characteristics of the production function that determine the need to provide incentives. Second are the characteristics of the agent, in particular whether he is risk averse or subject to limited liability and whether he has free access to credit markets. Third are transaction costs.

Below, I use this framework to identify the variables that are likely to affect the choice of contractual structure within each dimension in practice.

A. Contract Duration: Long Term vs Short Term.

Long term contracts have three main advantages over short term contracts. First they give the agent a stake in future output and hence provide incentives for non observable investment. This is of crucial importance in agriculture because tasks such as tree maintenance and careful application of fertilizers and pesticides have a strong effect on future output. Other things equal, this implies long term contracts are more likely to be used when, due to the characteristics of the crop, investment is important for productivity, as is the case for trees as opposed to annual crops.⁶

Second, long term contracts can be used to smooth consumption and reduce the risk borne by the agent when he has no access to credit.⁷ If risk aversion is decreasing in wealth and poorer tenants are less likely to have access to credit, this implies long

⁵The principal and the agent can, in principle, agree to other output shares. In practice, however, all sample contracts prescribe a 50-50 split.

 $^{^6}$ See Bardhan [1984], Banerjee et al [2002], Bose [1993] for specific applications to tenancy contracts.

⁷Chiappori et al [1994] and Rogerson [1985] analyze the case of repeated moral hazard when the agent has no access to credit markets. In this context the optimal long term contract generally exhibits "memory", i.e. payments in each period are a function of past performance. Note that if the agent has access to credit markets the outcome of a long term contract can be replicated by a sequence of spot contracts and this rationale for long term commitment disappears. See Fudenberg et al [1990] and Malcomson and Spinnewyn [1988].

term contracts should be more likely to be offered to poorer tenants.⁸

Third, long term contracts entail lower transaction costs because they have to be agreed upon less frequently. An implication is that long term contracts should be more common when the opportunity cost of time of the involved parties is high.

Long term contracts however entail a cost since commitment implies that the principal forsakes eviction threats, which could otherwise be used to elicit effort for current production. The threat of eviction in case of failure is an effective incentive mechanism when the agent's utility from the contract is higher than his reservation utility. Since this is more likely to occur when the agent is poor or has a low outside option, this implies poor tenants should be less likely to be offered a long term contract.⁹

Moreover, if the principal commits to a long term agreement, she gives up the possibility to adjust the terms of the contract to suit changes in the environment. In particular, the landlord gives up the option of cultivating the land directly for the duration of the contract and the contract reduces the resale value of the land if the buyer is bound to honor the existing tenancy agreement. The opportunity cost in terms of loss of flexibility is higher for landowners who might want or need to resume direct cultivation, implying that these should be more likely to offer short term contracts.

B. Contract Type: Fixed Rent vs Sharecropping.

Compared to share contracts, fixed rent contracts give the agent stronger incentives to exert non observable effort since under fixed rent he gets the full marginal benefit of his effort whereas under sharecropping he only gets a share. ¹⁰Other things equal, fixed rent contracts should therefore be chosen when the moral hazard problem is more severe, for instance because the cultivated crop is very sensitive to effort.

⁸See Bardhan [1983] and Fudenberg *et al.* [1990]. The latter also note that this prediction is in contrast with evidence from firms, since, compared to workers, managers are more likely to be offered a long term contract.

⁹Note that eviction threats provide incentives for both current effort and investment as the latter increases output in the next period and hence the probability of retaining the job in the period after next. See Banerjee et al [2002]; Banerjee and Ghatak [2003] and Dutta et al [1989]

¹⁰Singh [1989], Dutta et al [1989] and Otsuka et al. [1992] provide excellent surveys of the theoretical literature.

In addition, if the contract is fixed rent the principal does not need to monitor the division of output to make sure she is effectively getting the contracted share. Fixed rent contracts are therefore particularly well suited for landlords whose opportunity cost of time is high.

Fixed rent contracts can however be suboptimal, from the principal's point of view, for the following reasons. First, under fixed rent the agent bears all production risk. If the agent is risk averse, the principal might prefer sharecropping contracts as these strike a compromise between incentives and insurance. Share contracts should then be more common when the crop is risky and, if risk aversion decreases with wealth, when the tenant is poor.¹¹

Second, if the agent is subject to limited liability he might not be able to afford to pay rent in case of low output. The principal might then prefer to charge state contingent payments, in other words, offer a share contract.¹² Since the limited liability constraint is more likely to bind for poor tenants and tenants with low outside option, these should be more likely to be hired under share contracts. Share contracts should also be more likely when the spread between output in different states of nature is high.¹³

Finally, if production depends on both non observable effort and non-observable investment, share contracts might be preferred because fixed rent provide too much incentive for effort at the expense of investment. A similar argument can be made if it is the case that the tenant can increase current production at the expense of future production by overworking the land. This implies that share contracts should

¹¹See Stiglitz [1974].

¹²See Shetty [1988], Dutta et al [1989], Mookherjee [1997], Banerjee et al [2002]. Basu [1992] and Ghatak and Pandey [2000] also allow the tenant to choose the riskiness of the production technique. They show that limited liability leads to an inefficient outcome because it makes the tenant choose techniques that are too risky. In this setting sharecropping contracts might be preferred because they mitigate the incentive to choose risky projects. In the context analysed in this paper, however, considerations of this sort are not relevant as tenants have little discretion over production techniques.

¹³To see this, assume there are only two states of nature, good and bad, and that output in the bad state is zero. In the bad state, the maximum rent the tenant can pay is equal to his wealth minus subsistence consumption. As the output in the good state increases the rent the landlord wants to charge increases as well. For a given level of tenant's wealth, the limited liability constraint is therefore more likely to bind for crops that have a higher return spread. Banerjee *et al* [2002] and Mookherjee [1997] provide a formal analysis.

be observed when crop characteristics are such that multitasking issues are relevant as in the case of trees compared to annual crops.¹⁴

C. Summary.

Table 1 presents a summary of the discussion above. The table lists the main assumptions about the characteristics of the environment and their consequences for the choice of contract duration and type. These are then mapped into observable variables and implications are drawn. The variables and implications have been selected keeping in mind the particular context of nineteenth century rural Sicily and the available data, hence they do not constitute an exhaustive list.

The table highlights two important issues. First, different assumptions lead to similar implications for the effect of one variable on one contractual dimension different on the other dimension. Information on both dimensions can then be used to assess which considerations prevail. For instance, both limited liability and risk sharing considerations imply that sharecropping contracts should be used for poor tenants and risky crops. However, limited liability (with risk neutrality) implies that poor tenants should be offered short term contracts while risk sharing points to the opposite. Which effect prevails is ultimately an empirical question.

Second, data on both dimensions of contractual structure allow a better understanding of the evidence. For instance, multitasking considerations would suggest that since fixed rent contracts provide too much incentive for production effort at the expense of investment effort, sharecropping should be used for crops that are more sensitive to investment effort, namely trees. Contract length, however, can be used to provide investment incentives directly, thus weakening this rationale for sharecropping.

3 Data Description and Methodology

A. Data Description: Historical Context and Main Variables.

I use information on 705 tenancy contracts written in the district of Siracusa,

¹⁴See Holmstrom and Milgrom [1991], Allen and Lueck [1996], Ackerberg and Botticini [2000] and Dubois [2002] for discussion and evidence on multitasking.

Italy, between 1870 and 1880. Agriculture was the most important economic activity at the time, employing the majority of the work force. Tenancy agreements were common since land was unevenly distributed and rarely cultivated by the owners.¹⁵ Each contract is a legally binding agreement between a landlord, who owns the plot, and a tenant who is hired to cultivate it. Contracts were written by a notary public, following the instructions of the parties, and signed by these in his presence.¹⁶

The time period is chosen to match with an extensive descriptive literature on Sicilian agriculture. In 1881 the Italian Parliament published a detailed survey on the economic and social structure of the agricultural sector in different regions of the country [Inchiesta Iacini, 1881]. A similar survey was also carried out in 1911 [Inchiesta Parlamentare, 1911]. Both surveys describe Sicilian agriculture in great detail and contain information that is relevant for the present work.

Compensation Scheme.

Each contract specifies the payment from the tenant to the landlord, which can either be a share of the output, a fixed payment (either monetary or in kind) or a combination of both. Most contracts (85% of the sample) are of the fixed rent type, that is the tenant retains all the output and pays a fixed amount to the landlord at the end of every year. The remaining 15% of contracts are of the sharecropping type with share equal to one half.¹⁷

¹⁵In Sicily, feudalism was officially abrogated in 1812. Feudal fiefs were subsequently divided but most landholdings remained quite large and in the hands of the aristocracy or rich *burgeoise*, who typically rented out. Farmers who rented in were landless or owned small plots, insufficient for subsistence (Inchiesta Iacini [1881], Inchiesta Parlamentare [1911]).

¹⁶Compared to verbal agreements or contracts written privately by the two parties, contracts written by notary publics had the status of "public" documents, which made them safer for both parties. First, public contracts were binding for third parties implying that, for instance, if the landlord were to sell the land the buyer had to honor the existing tenancy agreement. Also, in case of sale, the tenant would give up the right to demand compensation. Finally, since most tenants were illiterate, notary contracts make sure that the landlord effectively wrote what was verbally agreed upon. See Codice Civile per il Regno d'Italia (1865), no. 1597, 1601.

¹⁷The variation in contract type is much lower than what theory would predict but consistent with observations from many other rural contexts. See Ackerberg and Botticini [2002], Dubois [2002], Laffont and Matoussi [1995], Young and Burke [2001].

¹⁸The existing law ruled that, unless specified otherwise in the sharecropping contract, the tenant was entitled to one half of the output. In addition, the tenant was supposed to provide draft animals, tools, working capital and finance all "ordinary" cultivation expenditures. The landlord was supposed to replace plants, if needed, and to finance "extraordinary expenditures". Finally,

Interestingly, the fact that under sharecropping the tenant had the incentive to cheat on the division of output was acknowledged by the judicial authorities at the time. To protect the landlord from tenants' opportunism, the law ruled that the tenant could harvest the crop only after giving the landlord notice.¹⁹

Contract Duration.

Contracts make precise the duration of the agreement. Figure 1 shows that although duration ranges from one to ten years, most contracts in the sample are either 1 or 4 years long. The fact that the unit of measure of duration is years as opposed to, say, months or days, is due to the fact that all crops in the sample give yearly yields. The concentration on two values (one and four) is more surprising. This might be due to the same reasons that limit the variation of the output share and is also consistent with evidence from other studies of contract duration.²⁰

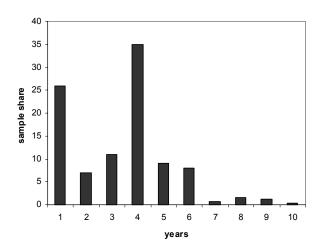


Figure 1: Contract Duration

Given the distribution of the duration variable and for ease of exposition, the analysis focuses on the distinction between short and long term contracts. The former

the landlord and the tenant each had to provide half of the seeds. See Codice Civile per il Regno d'Italia (1865) no.1654, 1655, 1656, 1657, 1658, 1661. See also Pacifici-Mazzoni and Venzi (1921) p. 343, 352, 353, 355, 356, 365, 366.

¹⁹Codice Civile per il Regno d'Italia (1865) no.1660. See also Pacifici-Mazzoni and Venzi (1921) p. 367.

²⁰For instance, Brickley et al [2003] find that most franchise contracts are 5 or 10 years long.

includes one year contracts that give the tenant no stake in future production; the latter includes contracts that are longer than one year and therefore provide some incentives for non observable investment effort. As shown in section 5, using all information on duration does not offer additional insights.

It is important to note that the length variable measures the duration of the contract, which does not necessarily coincide with the duration of the relationship between the landlord and the tenant. This could indeed be much longer if the same parties were to renew the agreement every time it expires.

It is then key to assess whether the duration of the contract effectively conveys information on investment incentives, or whether, because of frequent renewals, short term contracts are practically equivalent to long term contracts.

If the tenant expects to leave the plot at the end of the year with positive probability, a sequence of short term contracts is not equivalent to a long term contract with regards to investment incentives because a positive probability of non-renewal is effectively a tax on the return of the tenant's investment effort.

In this context there are two reasons to believe that the ex-ante probability of non-renewal is indeed positive. First, the wording of the contracts makes clear that the tenant is required to leave the land at the end of the lease.²¹ Second, since signing new contracts is costly both in terms of time and because the notary public charges a fixed fee for his services, the *ex-ante* probability of non-renewal must be positive, otherwise parties could save on renegotiation costs by signing a long term agreement.²²

²¹By law, each contract terminates by law on the last day of agreed lease period. If the duration is not specified the contract is intended to be expire after the first harvest. For both sharecropping and fixed rent contracts, if the tenant remains on the plot with the consent of the landlord the contract is extended until the next harvest but it loses the status of public document unless it is formally renewed in the presence of a notary public. See Codice Civile per il Regno d'Italia (1865), no. 1591, 1593, 1622, 1623, 1624, 1651. See also Pacifici-Mazzoni and Venzi (1921), p.202-207, 273-275, 375-377.

²²In addition, rent reduction rules made long term contracts more convenient for the landlord. The law indeed prescribes that if, due to circumstances beyond the tenant's control, more than half of the harvest got destroyed, the tenant has the right to demand rent reduction in proportion to the loss if the contract is one year long. If the contract is long term, however, the tenant loses the right to demand rent reduction if the loss in one year is compensated by rich harvests either in past or in future years. See Codice Civile per il Regno d'Italia (1865), no. 1617, 1618; Pacifici-Mazzoni and Venzi (1921), p. 256-272.

It is important to note that the *ex-ante* probability of non-renewal can be positive even if the same landlord-tenant pair sign a short term contract year after year. Indeed, when eviction threats are used as an incentive mechanism, the *ex-post* probability of renewal is high precisely because the *ex-ante* threat of non-renewal provides effort incentives leading to high productivity. As argued in section two, giving up eviction threats is part of what makes long term contracts costly.

In addition to the information on the terms of the agreement, each document also contains information on the type of crop, on the wealth of the tenant and on landlord's characteristics. The remainder of this section describes these in detail.

Crop Type.

Crops in the sample are cereals, olive, vines, citrus and fruit trees.²³ With a few exceptions, each contract regulates the cultivation of one crop only. In most cases where annual and tree crops are grown in the same plot, as is sometimes the case with wheat and olive, trees are excluded from the agreement.²⁴ The typical contract also contains clauses to forbid tenants to change crops or to plant other crops in addition to the existing ones.

Detailed information on crop characteristics can be found in Inchiesta Iacini [1881] and Inchiesta Parlamentare [1911]. According to these, tree crops were much more sensitive to non observable investment effort than the annual crops in the sample, which is consistent with intuition and with evidence from other times and places.²⁵

Fertilizers and/or manure were seldom employed in wheat and barley fields. Instead, the land would be left fallow every three or five years to restore fertility. Fallow land is obviously observable and hence contractible, suggesting that only a few non observable investment tasks (e.g. deep ploughing and weeding) were left to the dis-

²³Cereals are wheat and barley, whose cultivation techniques were very similar. Eighty-two percent of cereals contracts are for wheat. Fruit trees include many varieties such as cherries, pears, peaches, apricots and almonds.

²⁴This was possible because trees were generally grown on one side of the plot, which the tenant was asked to ignore. Contracts typically contain a detailed description of the location and number of trees to prevent the tenant from cutting them to sell the wood.

²⁵Ackerberg and Botticini [2000 and 2002] argue that vines were more sensitive to investment effort than cereals in Renaissance Tuscany. Holmstrom and Milgrom [1991] make a similar point about vines in contemporary California.

cretion of the tenant. Non observable investment effort was much more important for vines and citrus trees, which were very sensitive to the timing and dosage of fertilizers and pesticides. Both crops needed regular hoeing (4/5 times a year) and pruning, ²⁶ citrus trees were also very sensitive to irrigation timing. Olive trees also needed regular pruning, careful harvesting, tilling and fertilizing but were apparently more resistant than either vines or citrus.²⁷

While it is difficult to measure riskiness and spread of each crop precisely, the qualitative evidence indicates a clear ranking. Vines and citrus trees were the riskier crops, olive trees were somewhat less risky and cereals the safest. The ranking in terms of spread between the good and the failure state is similar, namely vines and citrus had the highest. For instance, Inchiesta Iacini [1881] reports that net revenues per hectare under "normal condition", that is in the success state, ranged between L.50 and L.150 for wheat, between L.300 and L.800 for vines and between L.500 and L.1300 for citrus trees.²⁸

In the empirical analysis I group crops in two ways. The most conservative choice, used for the main specification, exploits the natural difference in life span as a measure of investment sensitivity, thus I group all tree crops together and compare them to annual crops. Alternatively I use the available evidence on the difference between vines and citrus on the one hand and olives and fruit on the other to form three crop groups: annual (cereals), low-maintenance/low-risk trees (olive and fruit) and high-maintenance/high-risk trees (vines and citrus).

Table 2B shows the frequency of the different type of contracts by crop type. Contractual structure clearly varies by crop: 89% of annual crop fields are cultivated under fixed rent, with a predominance of short (53%) over long term (36%) contracts. Sharecropping is much more likely for tree crops, especially for vines and citrus. The

²⁶To avoid excess pruning motivated by the resale value of the wood, contracts typically established that the pruned woods belonged to the landlord. In some cases pruning was performed by other workers under the direct supervision of the landlord.

²⁷Olive yields were particularly sensitive to the harvesting method employed the year before. The quickest system, "abbacchiatura", consisted in shaking the tree until all the olives fell. This system had the serious drawback of destroying many of the buds, thereby reducing the following year's production. See Inchiesta Iacini [1881].

²⁸No information on net revenue of either olive or other fruit trees is reported, possibly because these were generally grown for personal consumption rather than commercialization.

difference in contract length between annual and tree crops is striking: while about 60% of the contracts for annual crops are one year long, the percentage falls to 8% for all tree crops and only 3% for vines and citrus.²⁹

Tenants' Wealth.

In nineteen century rural Sicily, formal credit markets were seriously underdeveloped and accessible only to wealthy landowners since lenders required strong guarantees. Poor farmers relied on informal lenders and on their landlords for working capital loans. The fact that credit and insurance markets were highly imperfect is especially important because it suggests that risk sharing and limited liability issues, both of which make incentive provision costly, are relevant in this context.

Each contract in the sample specifies the social class of the tenant after mentioning his name. Social class can be reasonably used to proxy wealth and both the Inchiesta Iacini [1881] and the Inchiesta Parlamentare [1911] report a clear ranking of rural social classes according to the wealth of their members.

Tenants in the sample belong to one of three social classes. The lowest class was made of *villici*, poor farmers who owned "only the strength in their arms" [Inchiesta Iacini 1881]; the second lowest were *contadini* or *coloni*, farmers who owned a mule and/or a small house and possibly a small plot of land by the house; the wealthiest class were *possidenti*, that is tenants who owned land of their own.³⁰ I refer to tenants in the three classes as *poor*, *middle class* and *rich* respectively. The sample shares are 34%, 40% and 26%.

Table 2C shows the frequency of the different type of contracts by tenant class. The table reveals that the frequency of long term-fixed rent contracts increases steadily with wealth while the frequency of short term-fixed rent contracts declines dramatically as wealth increases. Also, sharecropping contracts are twice as likely for poor, compared to rich, tenants.

²⁹Average duration is 2.3 for cereals, 3.85 for olives and fruit trees, 4.15 for vines and citrus trees. ³⁰A small number of tenants (5% of the sample) belonged to the class of *massari*, that is wealthy farmers who owned draft animals, a house and some plots of land. For simplicity, these have been grouped with the wealthiest class of *possidenti*. Moreover, *industriosi*, i.e. artisans, whose wealth, according to Damiani (1881), was comparable to *contadini's* have been included in that group. Results are robust to alternative definitions.

Social class improves over existing wealth measures because it can be taken to be exogenous to the extent that social mobility is low. Such was the case in 19th century Sicily, where, according to Inchiesta Iacini [1881] and Inchiesta Parlamentare [1911], social class was generally determined at birth. However, since social class is quite coarsely defined, the estimate of the wealth effect might be biased downward.

Limited liability models predict that contract choice depends on the tenant's outside option, in addition to her/his wealth. Contracts do not contain this type of information but Inchiesta Iacini [1881] reports data on the daily wage for rural workers in different towns. I use this as a proxy for the tenants' outside option. The average wage was L.1.38 with a standard deviation of .27.³¹

Landlord's Characteristics.

Contracts contain information on the gender of the landlord, on whether her legal residence was in the same town where the plot was located, and on whether she belonged to the aristocracy. These variables proxy for the landlord's participation in the agricultural business and hence for monitoring, transaction and flexibility costs.

Due to social norms, female landlords were not likely to be directly involved in cultivation and, due to the fact that they had to travel from a different town, landlords who resided away from the plot were also less likely to participate to agricultural decisions. Landlords who belonged to the aristocracy were also less likely to be directly involved in agriculture. Female, absentee and aristocratic landlords faced a higher opportunity cost of time and higher monitoring and renegotiation costs. To the extent that they were less likely to either need or want to resume direct cultivation, these landlords were also likely to value flexibility less.

Table 2D shows the frequency of the different type of contracts by landlord's characteristics. Aristocratic and female landlords were clearly different from the average landlord in the sample as they were much more likely to offer long term/fixed rent contracts (88% vs 62% in the overall sample). Landlords whose legal residence was in a different town also seem very different from the average as they are, surprisingly, more likely to offer short term contracts.

³¹Italian Liras ca 1881.

Note that "legal residence" identifies the town where the landlord is registered with the records' office rather than the town she lived in. Although most of the times these should coincide, landlords would not change their records if they moved temporarily to another town. To the extent that this happened, "legal residence" is a noisy measure of actual residence.

B. Methodology.

The analysis focuses on the determinants of the choice between long term and short term contracts and between fixed rent and sharecropping contracts. To begin with, contracts are classified as long term (short term) if they last more (less) than 1 year. The classification is motivated by the consideration that, in contrast to longer contracts, one year contracts give the tenant no stake in future production. In section 5, I extend the framework to allow contract duration to take multiple values.

The landlord chooses the length of the contract and the compensation scheme to maximize her payoff for given tenant's and crop's characteristics.³² As discussed above, under fixed rent contracts the tenant's output share is equal to one while under sharecropping contracts the share is one half. The two first order conditions of the landlord's maximization problem yield the optimal length (l^*) and output share (f^*) as a function of each other and of the exogenous variables. Assuming linearity, the model is;

$$\begin{cases} l^* = \alpha_l f^* + X' \beta_l + \epsilon_l \\ f^* = \alpha_f l^* + X' \beta_f + \epsilon_f \end{cases}$$
 (1)

where X' is the vector of observable tenant, landlord and crop characteristics while ϵ_l , ϵ_f capture the effect of variables that affect the landlord's choice but are not observed by the econometrician. Variables in X include the tenant's social class, the landlord's gender, social class and town of residence and the type of crop cultivated

³²In line with most literature, the landlord is assumed to have all the bargaining power and matching between landlords and tenants is assumed to be random. The first assumption is appropriate given the abundance of labor relative to land in the context under study. Random matching is discussed in section 5, below.

on the rented plot. Variables in ϵ_l , ϵ_f include for instance soil quality and the tenant's degree of risk aversion.

The data does not contain information on l^* and f^* . Instead, we observe two discrete variables l and f; where l equals one when the contract is long term and zero otherwise and f equals one when the contract is fixed rent and zero otherwise. If l^* and f^* are global maxima, the landlord will choose a long term contract if the optimal length is above a threshold \bar{l} and similarly choose a fixed rent contract if the optimal share is above a given threshold \bar{f} . The decision rule then is;

$$\begin{cases} l = 1 & if \quad l^* > \bar{l} \\ l = 0 & if \quad l^* \le \bar{l} \end{cases} \text{ and } \begin{cases} f = 1 & if \quad f^* > \bar{f} \\ f = 0 & if \quad f^* \le \bar{f} \end{cases}$$
 (2)

The discussion in section 2 makes clear that none of variables in X can be reasonably excluded from either the length or the type equation in (1). In addition, neither contract law nor other exogenous factors that could affect contractual structure exhibit geographical or time variation in the sample. Following standard practice in the empirical contract literature,³³ I therefore estimate the reduced form of (1);

$$\begin{cases} l^* = \gamma_l + X' \pi_l + \nu_l, & l = 1 & \text{if } l^* > 0, & l = 0 & \text{otherwise} \\ f^* = \gamma_f + X' \pi_f + \nu_f, & f = 1 & \text{if } f^* > 0, & f = 0 & \text{otherwise} \end{cases}$$
 (3)

Where γ_l, γ_f are constants, X is the vector of tenant, landlord and production function characteristics. To take into account that some unobserved determinants might be common to both equations in (2), I assume that the disturbances ν_l and ν_f are jointly normally distributed with $E[\nu_l] = E[\nu_f] = 0$, $Var[\nu_l] = Var[\nu_f] = 1$, $Cov[\nu_l, \nu_f] = \rho$ and estimate the system by bivariate probit.

In terms of the structural form parameters, the coefficient on variable k in equation i is equal to $\pi_i^k = \frac{\beta_i^k + \alpha_i \beta_j^k}{1 - \alpha_i a_j}$ for i, j = l, t. This implies that if $\alpha_i \neq 0$ the reduced form coefficients π_i capture both the direct effect of each variable on each dimension (β_i) and the indirect effect through the other dimension (β_j) .

The structural coefficient of variable k in equation i in (1) is then equal to $\beta_i^k = \pi_i^k - \alpha_i \pi_j^k$, which implies that if the indirect effects are not too large, namely $|\alpha_i| \leq \left|\frac{\pi_i^k}{\pi_j^k}\right|$, the structural coefficient will have the same sign as its reduced form counterpart.

³³See the numerous studies reviewed in Chiappori and Salanie [2003].

In addition to the coefficients vectors (π_l, π_t) , I evaluate the marginal effect of each right hand side variable on the marginal probability of observing a long term contract $\Pr(l=1)$, on the marginal probability of observing a fixed rent contract $\Pr(f=1)$ and on the joint probabilities of the four possible length and type combinations.

The marginal effect on the marginal probabilities measures the effect of one exogenous variable on one contractual dimension at the time. Given that exogenous variables are discrete the marginal effect is measured as the discrete change in probability when the exogenous variable goes from 0 to 1. For instance, the marginal effect of x^k on the probability of observing a long term contract is

$$\gamma_{l=1}^k = \Pr(l=1|x^k=1) - \Pr(l=1|x^k=0)$$

Probabilities are computed using the cumulative normal distribution evaluated at $x^k = 1$ or $x^k = 0$ and at the mean of the other right hand side variables.

The second type of marginal effects measures the effect of one exogenous variable on both dimensions at the same time, namely on the joint probability of observing a given combination of length and type. For instance, the marginal effect of x^k on the joint probability of observing a long term (l = 1) and fixed rent (f = 1) contract is

$$\gamma_{l=1,f=1}^k = \Pr(l=1,f=1|x^k=1) - \Pr(l=1,f=1|x^k=0).$$

where probabilities are computed using the bivariate cumulative normal distribution, evaluated at $x^k = 1$ or $x^k = 0$ and at the mean of the other right hand side variables.

Note that the marginal effect of x^k on one marginal probability, for instance the probability of observing a long term contract, is a combination of its marginal effects on the joint probability of observing a long-fixed contract and on the joint probability of observing a long-share contract. In particular, when x^k is discrete, the marginal effects on the joint probability and the marginal effect on the marginal probabilities are linked as follows;

$$\gamma_{f=z}^k = \gamma_{l=1,f=z}^k + \gamma_{l=0,f=z}^k$$
 and $\gamma_{l=v}^k = \gamma_{l=v,f=1}^k + \gamma_{l=v,f=0}^k$

where v and z can be either 0 or 1. The marginal effect of x^k on the probability of observing f = z is thus the sum of the marginal effect on the joint probability

of observing f = z and l = 1 and the marginal effect on the joint probability of observing f = z and l = 0.34 Analyzing the marginal effects on the joint probabilities is of interest because it allows to identify cases in which the marginal effect of x^k on one contractual dimension is zero because the variable's effect on the other dimension cancels it out.

Finally, the estimation rests on the assumption that the right hand side variables are not correlated with the error term. First, crop choice is assumed to be exogenous to contract type. Although both are variables of choice for the landowner, the assumption is supported by the fact that the life span of the sample trees is much longer than the typical contract duration. Vines and citrus trees have a productive life of at least thirty years, while olive trees can last over one hundred years. Contracts in the sample are typically one or four years long. In this sense it is safe to assume that the landlord chose the contract to fit the crop rather than *vice versa*. ³⁵

Second, it is assumed that the type of contract does not affect the social class of the tenant, which is supported by the fact that social mobility at the time was extremely low. Wealth estimates might however be biased if tenants and crops are endogenously matched; I address this issue in Section 5.

Third, it is assumed that contract type does not determine the landlord's characteristics either. Of these two are clearly predetermined (gender and aristocracy) and

$$\Pr(l = v, f = z) = (\Pr(l = v | f = z)) * \Pr(f = z)$$

the marginal effect of x^k on Pr(l=1, f=z) is equal to:

$$\gamma_{l=1,f=z}^{k} = [\Pr(l=1|f=z,x^{k}=1) * \Pr(f=z|x^{k}=1) - \Pr(l=1|f=z,x^{k}=0) * \Pr(f=z|x^{k}=0)]$$

while the marginal effect of x^{k} on $\Pr(l=0,f=z)$ is equal to:

$$\gamma_{l=0,f=z}^k = [\Pr(l=0|f=z,x^k=1) * \Pr(f=z|x^k=1) - \Pr(l=0|f=z,x^k=0) * \Pr(f=z|x^k=0)]$$
 Therefore,

$$\gamma_{l=1,f=z}^k + \gamma_{l=0,f=z}^k = \Pr(f=z|x^k=1) - \Pr(f=z|x^k=0) = \gamma_{f=z}^k$$

³⁴This can be shown using the link between joint and conditional probabilities. Given that

 $^{^{35}}$ Endogeneity is a much more serious concern when all crops are annual, and can therefore be chosen at the same time as the contract (Dubois 2002).

legal residence would only be affected if crop choice would cause the landlord to move indefinitely to another town. Even if the landlord's characteristics are predetermined, sample selection and measurement problem could still create a spurious link between contract type and the other independent variables. Section 5 discusses these issues in detail.

4 Empirical Analysis: Main Findings.

A. Basic Specification

Table 3 reports the estimates of model (2). Columns type (a) estimate the probability of observing a long term contract, columns type (b) estimate the probability of observing a fixed rent contract. Each pair of equations is estimated simultaneously as explained above.

For each of the independent variables, Table 3 reports the reduced form coefficients π_l, π_f , their standard errors and the marginal effects on the marginal probabilities of observing a long term and a fixed rent contract respectively.

Columns 1a and 1b analyze the effect of crop type, columns 2a and 2b add information on the tenant's social class, while columns 3a and 3b also include landlords' characteristics. Point estimates and significance levels do not change in the three specifications, suggesting low correlation among the right hand side variables.

Contracts were signed in eleven different towns (all in the district of Siracusa) and ten different years (1870-1880).³⁶ The effect of the variables of interest is thus identified from the variation both within and across towns. To the extent that town and year specific unobservables affect both contractual structure and the right hand side variables, the coefficients might be biased. Columns 4a and 4b include town and year dummies to address this concern. The results are qualitatively similar to those in columns 3a and 3b but, not surprisingly given sample size, coefficients are smaller and the estimated marginal effects are between one half and one quarter of their previous values. Further results, not reported for reasons of space, show that changes are mostly due to the inclusion of the town dummies while including year

³⁶The eleven towns are: Augusta, Buccheri, Buscemi, Carlentini, Ferla, Francofonte, Lentini, Noto, Pachino, Rosolini and Siracusa.

controls is of little consequence. Both town and year dummies are jointly significant in both equations. In what follows, I take the most conservative strategy and include town and year controls throughout. The effect of the variables of interested is thus identified from the variation within town and year.

Crop Type.

The coefficients on crop type indicate that trees make long term contracts more likely but have no effect on contract type.

Column 3a shows that tree cultivation increases the probability of observing a long term contract by .40, that is more than half of the sample mean (.74). When the estimate relies exclusively on within town and year variation (column 4a), the marginal effect is .14.

Columns 3b and 4b show that when town and year effects are not controlled for, there is some evidence that trees reduce the probability of observing a fixed rent contract (by .06, significant at the 10% level) however this effect loses significance when towns and years effects are included.

Tenant Wealth

Columns 3a and 4a, table 3, show that when the tenant belongs to the lowest social class the probability of observing a long term contract falls by .3 when town and year effects are not controlled for and by .07 when they are. The estimated marginal effect is between 40% and 10% of sample mean. When the tenant belongs to the middle class, the probability of observing a long term contract falls by .22 in column 3a but the effect loses significance when I control for town and year effects (column 4a).

Results in column 3b and 4b indicate that the tenant's social class is the most important determinant of contract type. The probability of observing a fixed rent contract falls by .15 if the tenant belongs to the lowest, compared to richest, class and by .08 if he belongs to the middle class. With town and year controls, column 4b, the estimated marginal effects are .08 and .05 respectively.

Landlord Characteristics

The probability of observing a long term contract is higher when the landlord is female or belongs to the aristocracy. The magnitude of these effects is similar (.17 vs .16 in column 3a and .05 vs .06 in column 4a) while the coefficient of landlord's residence is not precisely estimated.³⁷

Columns 3b and 4b reveal a similar pattern for contract type. The probability of observing a fixed rent contract is higher when the landlord is female or belongs to the aristocracy and the effects are similar in size (.09 vs .08 in column 3b and .03 in column 4b) while landlord's residence has a positive and significant effect only when town and year effects are not controlled for.

Summary and Interpretation

The theories reviewed in section 2 yield three unambiguous predictions on the sign of the coefficients of the structural model (β_l, β_f) :

- (1) $\beta_l^{tree} \geq 0$; that is long term contracts should be more likely for trees as these are more sensitive to investment;
- (2) $\beta_f^{poor} \leq \beta_f^{middle} \leq 0$; the probability of observing a fixed rent contract should be decreasing in tenants' wealth, either because of risk sharing or limited liability;
- (3) $\beta_i^{absentee} \geq 0$, $\beta_i^{female} \geq 0$, $\beta_i^{aristocrat} \geq 0$ for i = l, f; that is, landlords who are less likely to be directly involved in cultivation should be more likely to offer long term fixed rent contracts to minimize transaction and monitoring costs.

Predictions on the other parameters, for instance the effect of tenants' wealth on contract length, are ambiguous as different effects pull in different directions.

The structural coefficients are a combination of the reduced form coefficients estimated in table 3. In particular, $\beta_i^k = \pi_i^k - \alpha_i \pi_j^k$, which implies to the extent that the direct effect of variable k on contractual dimension i is larger than its indirect effect through the other dimension, the sign of the structural coefficients are the same as the sign of the reduced form coefficients reported in table 3.

³⁷The coefficient is negative when cross-town and cross-year variation is not controlled for and zero when it is. Below, I find that this effect is not very robust and that it mostly derives from the fact that these landlords are less likely to offer long term *share* contracts rather than long term *per se*. As discussed above, measurement error in this variable is likely to be high because "legal residence" does not necessarily coincide with the town the landlord lives in.

In terms of the parameter of the model this requires $|\alpha_i| \leq \left|\frac{\pi_i^k}{\pi_j^k}\right|$, that is the effect of contract length on contract type and vice-versa should not be too large. If this condition is satisfied, the results in table 3 are consistent with the theoretical predictions. For instance, the estimated coefficients of tree in columns (4a) and (4b) indicate that as long as $\alpha_l > -22.15$, $\beta_l^{tree} \geq 0$ namely, long term contracts are more likely when the marginal benefit of investment is highest. Similarly, the estimates suggest that if $\alpha_t \leq 1.45$ poor tenants are less likely to get fixed rent contracts in line with both the risk sharing and the limited liability models. Finally, landlords who face higher transaction and monitoring costs are more likely to offer long term and fixed rent contracts.

Table 3 also reports the estimated correlation coefficient between the disturbance terms of the contract length and contract type equations. The estimate is positive (.08) indicating that omitted factors push for high powered incentives in both contractual dimensions but the Wald test cannot reject the null hypothesis of zero correlation. Since the correlation coefficient between the disturbance terms in the reduced form equations measures the correlation of the observed outcomes after controlling for exogenous variables, this finding does not contradict the idea that the choices over the two contractual dimensions are correlated in the underlying model.³⁸

B. Marginal Effects on the Joint Probabilities.

Table 4 reports the marginal effects on the joint probabilities of the four combinations, namely long term-fixed rent, long term-share, short term-fixed rent and short term-share. For ease of comparison the table also reports the marginal effects on the marginal probabilities from columns (4a) and (4b) in table 3 above.

Three findings are noteworthy. First, the effect of tenants' wealth on the probability of observing a long term contract depends on whether the contract is fixed rent or sharecropping. Column (1) indicates that, compared to rich tenants, poor and middle class tenants are significantly less likely to be offered a long term/fixed rent contract

$$Cov(\nu_l, \nu_f) = \frac{1}{(1 - \alpha_l \alpha_f)^2} [\alpha_f Var(\epsilon_l) + \alpha_l Var(\epsilon_f) + (1 + \alpha_l \alpha_f) Cov(\epsilon_l, \epsilon_f)]$$

that is $Cov(\nu_l, \nu_f) = 0$ does not imply $Cov(\epsilon_l, \epsilon_f) = 0$.

³⁸Note that since $\nu_i = \frac{\epsilon_i + \alpha_i \epsilon_j}{1 - \alpha_i \alpha_j}$, where ϵ_i are the disturbances of the structural equations,

(by .14 and .07). Column (2) shows that, at the same time, poor and middle class tenants are significantly more likely to be offered a long term/share contract (by .07 and .05). For poor tenants the first effect dominates, making long term contracts less likely to be observed when the tenant belongs to the lowest social class. For middle class tenants the two effects cancel out, so that the variable has no significant effect on the overall probability of observing a long term contract.

Second, the effect of trees on the probability of observing a fixed rent contract depends on whether this is long or short term. Column (1) suggests that, compared to annual crops, trees significantly increase the probability of a long term fixed rent contract (by .13) while column (3) shows that trees significantly decrease the probability of a short term fixed rent contract (by .14). The two effects cancel out and, as seen in column 6, trees do not significantly affect the overall probability of observing a fixed rent contract.

Third, female and aristocratic landlords are significantly more likely to offer long term fixed rent contracts and significantly less likely to offer any of the other three combinations. Landlords whose legal residence is in a different town are less likely to offer long term share contracts.

Interestingly, the table shows that using information on one dimension only can mislead the interpretation of the evidence if, as is sometimes the case here, the effect of one variable on one contractual dimension is nil because the variable's effect on the other dimension cancels it out.

Overall the results suggest that crop type primarily drives the choice between long and short term while tenant's wealth appears to be the main determinant of the choice between fixed rent and share. Landlords who face higher monitoring and renegotiation costs choose long term coupled with fixed rent.

C. Interactions.

Table 5 augments the model by allowing the effect of tenant class to depend on crop type and *vice-versa*. Three findings emerge.

First, the fact that the coefficient of *poor* is negative and significant in column (1a) indicates that rich and middle class tenants are more likely to get long term

contracts regardless of crop type.

Second, the fact that, in column (1a), the interaction term poor *tree is positive, significant and equal in magnitude but opposite in sign to the poor variable, indicates that when they cultivate trees poor tenants are as likely as other tenants to be offered a long term contract. In other words, poor tenants are more likely to get long term contracts only if they cultivate trees.

Finally, the fact that *poor* and *middle class* by themselves do not have a significant effect on contract type (column 1b) while they are negative and significant when interacted with *tree* suggests that tenants belonging to the two lower classes are as likely as rich tenants to get a fixed rent contract when they cultivate annual crops, whereas they are significantly more likely to get a share contract when they cultivate trees. This finding also speaks to the relationship between crop type and contract type. Trees are indeed, and as expected, more likely to be cultivated under share agreements but only if the tenant is not rich.

Overall the balance of evidence indicates that investment incentives (via long term contracts) are always provided for tree crops but not for annual crops and poor tenants. High powered incentives for both investment and effort are always offered when the tenant is rich, and when the landlord faces high monitoring and renegotiation costs.

To the extent that the structural coefficients have the same sign as their reduced form counterpart, the findings suggest that risk sharing considerations do not prevail in the choice of contract length. Poor tenants, who would need insurance against unemployment risk and income fluctuations the most, are actually less likely to be offered a long term contract.

Finally, that share contracts are more likely to be chosen for trees but only when the tenant is poor suggests that, in this context, multitasking considerations do not play a major role in determining the choice between fixed rent and share contracts. If they did, we would expect trees to be cultivated under share contracts regardless of the wealth of the tenant.

5 Empirical Analysis: Extensions.

Alternative Definition of Contract Length.

The analysis of contract length as a dichotomous variable highlighted the difference between one-year contracts that give the tenant no stake in future production, and longer contracts that, instead, make the tenant's pay conditional on future performance.

Since sample contracts are between one and ten years long, the data also allows to analyze the choice among all the different duration outcomes. The exercise is of interest in itself and because it allows to establish whether the results of the previous section are due to the particular classification of length employed there.

To capture the fact that the choice of length is discrete, I estimate the following model;

$$\begin{cases} d^* = X' \pi_d + \nu_d, & d = j & \text{if} \quad k_{j-1} < d^* < k_j, \quad j = 1..10 & k_0 = -\infty; k_{10} = \infty \\ f^* = X' \pi_f + \nu_f, & f = 1 & \text{if} \quad f^* > 0, & 0 & \text{otherwise} \end{cases}$$
(4)

As in the previous section, I assume that the disturbances ν_d and ν_f are jointly normally distributed with $E[\nu_d] = E[\nu_f] = 0$, $Var[\nu_d] = Var[\nu_f] = 1$ and $Cov[\nu_d, \nu_f] = \eta$ to take into account that some unobserved determinants might be common to both equations. I then estimate the two equations in (4) jointly by full information maximum likelihood.

Table 6 presents the findings for both the basic specification and the model with interactions. Columns 1a and 2a report findings on duration, columns 1b and 2b on type. The corresponding results for the dichotomous length variable were presented in tables 3 and 5.

The estimated coefficients in the duration equations suggest that, in line with the previous findings, one year contracts are less likely to be used for trees, by landlords who do not belong to the aristocracy and for rich tenants. In contrast to previous findings, however, the landlord's gender is not a significant determinant of contract length.

To illustrate the effect of the exogenous variables on the choice of duration between

one and ten, Table 7 reports the marginal effect of each variable on the probability of observing each of the ten outcomes. Since all right hand side variables are discrete, the marginal effect of x^k on the probability of d = r, for r = 1...10 is;

$$\gamma_{d=r}^{k} = \Pr(d = r | x^{k} = 1) - \Pr(d = r | x^{k} = 0)$$

Three findings are noteworthy.

First, the comparison of these results to the ones for the dicothomous length variable used in tables 3 and 5 reveals that the estimates are identical in sign and very close in magnitude. To be precise, the marginal effects on the probability of observing a one year contract in table 7, are equivalent to the marginal effects on the probability of observing a short term contract when the length variable is dichotomous as in tables 3 and 5. Using all information on contract length does not therefore alter the previous findings.

Second, the marginal effects are by far larger for two most common outcomes, namely one and four years, compared to all the other alternatives, suggesting that the previous results on the choice between short and long term contracts were effectively driven by the comparison between one and four year long contracts.

Third, table 7 shows that the marginal effects on the probability of observing a one, two or three year long contract have the same sign and that the sign of the marginal effects switches for the probabilities of observing all durations larger or equal to four. For instance, tree cultivation increases the probability of observing contracts that are longer than three years and decreases the probability of observing contracts that are three years or shorter.

A plausible reason is that wheat was often cultivated in a two or three year rotation with fallow and legumes and the plot was rented out for the entire duration of the cycle. On a three year rotation, plots were divided into three parts and in each year a different part would be cultivated with wheat, one with legumes and the third left fallow. A three year contract would then be needed to complete the cycle on the plot. The two year rotation was similar but no part was left fallow.³⁹

³⁹Under the latter system the whole plot would generally be left fallow after two cycles, i.e. every fifth year.

In these cases it would take, respectively, three or two years for wheat to be harvested throughout the plot so that a three or two year contract would be effectively be "short term" by the definition above. Richer data would be needed to explore this issue further. For the purpose of this paper, it is reassuring to note that the marginal effects on the probabilities of observing a two or three year long contract are very small, so that the estimates of the dichotomous length variable are the same regardless of whether two and three year contracts are classified as long or as short term.

Finally, table 6 shows that allowing contract length to take ten values does not affect the findings on the determinants of contract type. The comparison of columns 1b and 2b (Table 6) to columns 4b (Table 3) and 1b (Table 5) reveals that the magnitude of the coefficients in the contract type equation is identical.

Alternative Definitions of Crop Type

As discussed in section 3, while their longer life span makes trees naturally more sensitive to investment than annual crops, trees do not constitute an homogeneous group. Agronomic evidence suggest that vines and citrus trees might be more sensitive both to production and investment effort than olives and fruit trees. In addition, vines and citrus trees are likely to be riskier and have a larger spread between good and bad state outcomes. Columns 1a and 1b, table 8, exploit this information and estimates model (1) keeping olives and fruit trees separate from vines and citrus trees.

Results in column 1a indicate that both types of trees increase the probability of observing a long term contract but the marginal effect of vines and citrus trees is about three times larger than the effect of olive and fruit trees (.13 compared to .04). The difference is statistically significant and consistent with vines and citrus trees being more sensitive to non observable investment. Column 1b shows that while vines and citrus trees decrease the probability of a fixed rent contract (and hence increase the probability of a share contract), olive and fruit trees increase it. However, neither effect is significant at conventional levels. Further analysis, not reported for reasons of space, reveals that, as with the previous definition of trees, share contracts are more likely to be observed when vines or citrus are cultivated and the tenant belongs to the two lowest classes. Finally, the comparison with the basic specification reported

in table 3, reveals that the estimated effects of all other variables are unchanged.

Although trees are often found in wheat fields, they are generally excluded from the contract. Only 15% of sample contracts for wheat also required the tenant to tend to trees as a secondary crop.⁴⁰ In columns 2a and 2b, table 8, the tree variable equals one both when trees are the primary crop and when they are a secondary crop in a wheat field. Coefficient estimates are generally unchanged except that the marginal effect of crop type on contract length doubles (.26 vs .13), which suggests that neglecting the information on secondary crops biases the coefficient of tree downwards.⁴¹

Tenants' Outside Opportunity.

If the tenant is subject to limited liability, the structure of the optimal contract depends on his outside option in addition to his wealth. In particular, fixed rent contracts should be more likely when the tenant's outside option is high. Columns 3a and 3b, table 8, include the daily wage for adult male rural workers in the local labor market as a proxy for the tenant's outside option. Data on wages are collected from interviews with the towns' mayors and reported in Inchiesta Iacini [1881] for each town.

To the extent that farmers are unwillingly or unable to move between towns the wage captures the differences in the value of the outside option among farmers residing in different towns.⁴² Given that wage does not vary within town, in columns 3a and 3b, table 8, I use information from the Inchiesta Iacini [1881] to group towns in regions with similar climate and soil characteristics.⁴³

⁴⁰There are 41 such contracts in the sample. Of these, 11 had olive trees as a secondary crop, 5 had vines and 25 had fruit trees.

⁴¹The estimated coefficient is the same if wheat contracts with trees as a secondary crop are dropped from the sample. The symmetric case (i.e. tree contracts with wheat as secondary crop) cannot be analysed because there are only 5 such contracts in the sample.

⁴²The wage data was collected only at one point in time during the second half of the decade but according to the Mayors the wage had hardly changed during the preceding *twenty* years. Since the effect of wages on contractual structure is identified from the variation across towns, the wage is a reasonable measure of the tenants' alternative option as long as the towns' relative ranking remained unchanged during the period analyzed here.

⁴³The three regions are: northeastern coastal (Augusta, Carlentini, Francofonte, Lentini, Siracusa), southeastern coastal (Noto, Pachino, Rosolini) and western interior (Buccheri, Buscemi, Ferla).

Caveats notwithstanding, results in columns 3a and 3b indicate that a fall in wage significantly reduces the probability of observing a fixed rent contract. The marginal effect of a decrease in wage by one standard deviation is equal to -.04, that is about one third of the effect of the tenant being middle class, instead of the rich, and about one fourth of the effect of tenant being poor, again instead of rich.

Endogenous Matching and Other Concerns.

The analysis so far has relied on the assumption that the landlord chooses the optimal contract for given characteristics of the tenant, or that, in other words, matching between landlords and tenants is random.

If this were not the case, however, the estimates would be biased. In their study of tenancy contracts in Renaissance Tuscany, Ackerberg and Botticini [2002] argue that endogenous matching of tenants and crops can lead to biases in both the crop and the risk aversion variable when the latter is not observed by the econometrician and a proxy needs to be used. They find strong evidence of matching, in particular that poorer tenants are more likely to farm vines instead of cereals. Controlling for matching changes their estimates considerably, most notably tenant's wealth becomes a significant determinant of the choice between sharecropping and fixed rent contracts.

To assess whether matching of tenants and crops is of concern in this setting I follow Ackerberg and Botticini [2002] and estimate the relationship between crop type and tenants wealth. Intuitively, if tenants of a given class systematically end up cultivating a given type of crop, tenants' class should be significantly correlated with crop type. Table 9 shows that in this sample there is no correlation between the two variables. Compared to rich tenants, tenants belonging to the two lowest classes are neither more likely nor less likely to cultivate a given type of crop. The result is robust to alternative crop classifications. In column (1) all tree crops are grouped together and compared to annual crops whereas in column (2) the riskier and more investment intensive crops (vines and citrus) are compared to annual and other tree crops. The coefficients of tenant's social class are not significant in either column, indicating that in this sample there is no evidence of matching between tenants and

crops.

Throughout the analysis, the reduced form coefficient of tenant wealth has been interpreted as a proxy for the cost of providing incentives, either because of risk aversion or limited liability. Tenants' wealth, however, might be capturing the effect of other tenant's unobservable characteristics. For instance, if wealth were a proxy for ability, less able, hence poorer, tenants could be offered short term contracts because landlords would not want to commit long term to a bad tenant. Also, richer, and hence more able, tenants could prefer to be residual claimants and get a fixed rent contract. The fact that wealth is coarsely proxied by social class combined with the fact that social mobility was very low, however, imply that the problem is much less severe in this context. Indeed, since social class was mostly determined by birth it is likely to be uncorrelated with innate individual traits, to the extent that these are randomly distributed across social classes.

Finally, since the landlord chooses both the crop and the contract, landlords unobserved characteristics might mislead the interpretation of the link between trees and contractual structure. To the extent that the available information on landlords' characteristics does not precisely capture the variables of interest, the results could be due to the residual variation in landlords' unobservable traits.

For instance, it could be argued that only landlords whose outside opportunity is very high would rent out trees instead of managing them personally. The observed correlation between trees and contract length could then be due to a selection bias if, at the same time, landlords with a higher outside option prefer to save on renegotiation costs and offer long term contracts. The fact that long term contracts are offered to rich tenants also when they cultivate annual crops however suggests that landlords' unobservable characteristics are not the sole determinant of contract length.

Information on the other contractual dimension sheds more light on this point. Based on the argument above, landlords whose outside opportunity is so high that they prefer to rent out their trees and to offer long term contracts to minimize renegotiation costs should, for the same reasons, avoid share contracts as these need the landlord to monitor the division of output and to sell her share of the agricultural produce. The empirical findings, in contrast, indicate that when the tenant is poor

trees are more likely to be cultivated under share agreements, which is *not* consistent with the idea that the link between tree and contract type is due to unobservable landlords' characteristics.

6 Conclusions.

This paper uses data on land tenancy agreements to present new evidence on the empirical determinants of contract form. The main novelty is the availability of evidence on contract length, an important and yet typically neglected dimension of contractual structure.

The evidence indicates that long term contracts are used for trees or when the tenant is rich, while poor tenants who cultivate annual crops are typically offered short term contracts. Evidence on the determinants of the compensation scheme suggests that fixed rent contracts are offered when the crop is annual or when the tenant is rich, while poor tenants who cultivate tree crops are typically offered sharecropping contracts. In addition, landlords who face higher renegotiation and monitoring costs are more likely to use long term and fixed rent contracts.

The findings provide evidence on the relevance of asymmetric information and on the factors that prevent the use of high powered incentives contracts. The fact that high powered investment incentives, via long term contracts, are always used for trees which are more sensitive to investment indeed suggests that asymmetric information plays an important role in contract design. Moreover, that poorer tenants are less likely to be offered high powered incentives is suggestive of the fact that credit market imperfections combined with either risk aversion or limited liability make incentive provision costly.

Importantly, the use of short term contracts and, relatedly, low powered investment incentives indicates that asymmetric information might effectively lead to less investment and lower productivity.

APPENDIX: DATA SOURCES.

Contracts.

Sample contracts are written by notaries public, who were required by law to store and bind all contracts they wrote each year. Upon retirement, contracts are stored at the Notary Archives and eventually transferred to the State Archives. Contracts in the sample were written by the following notaries and can be found at the State and Notary Archives in Siracusa, Italy:

| STATE ARCHIVES | | NOTARY ARCHIVES | |
|---------------------------|--------------|-----------------------|--------------|
| Notary Name | Sample Years | Notary Name | Sample Years |
| | | | |
| Baiona, Giovanni Battista | 1870-74 | Maxeo, Rosario | 1875-80 |
| Giarracca, Domenico | 1870-77 | Perricone, Alessandro | 1875-77 |
| Casaccio, Gaetano | 1870-72 | Tribulato, Antonino | 1876 |
| Di Giovanni, Ignazio | 1870-75 | Amico, Pietro | 1870-79 |
| Milito, Virgilio | 1870-72 | Tribulato Giuseppe | 1873-75 |
| Pancari, Vito | 1870-75 | Santuccio, Francesco | 1873-80 |
| Lenares, Giovanni | 1872-74 | Cultrera, Francesco | 1873-75 |
| Motta, Giuseppe | 1872-77 | Pisana, Pietro | 1875-76 |
| Pupillo, Giuseppe | 1870-75 | Sofia, Luigi | 1876-77 |
| Blasco, Francesco | 1870-71 | Sbano, Francesco | 1875-80 |
| Piccione, Rosario | 1870-76 | Tavana, Eustachio | 1876-80 |
| Terranova, Carmelo | 1870-77 | Nuzzo, Vincenzo | 1878-80 |
| Giardina, Gaetano | 1870-72 | Carbonaro, Raffaele | 1878-80 |
| | | Scalia, Gaetano | 1878-80 |
| | | Italia, Francesco | 1878-79 |
| | | Leone, Gaetano | 1880 |
| | | Zivillica, Alfonso | 1879-80 |

Town Variables.

From "Inchiesta Jacini: Atti della Giunta per l'inchiesta agraria e sulle condizioni della classe agricola 1881" vol XIII, parte I e II, tomo 1-5-rapporto di Abele Damiani per la Sicilia. Wage data: Book 2 A pages 303 and 311, Crop Revenues Data: Book 2 A pages 298 and 307.

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TABLE 1 DETERMINANTS OF CONTRACTUAL STRUCTURE: IDEAS FROM THEORY

| | ASSUMPTIONS | CONSEQUENCES FOR (| CONTRACTUAL STRUCTURE | RELATED VARIABLES | PREDICTIONS |
|--------------------------------------------------|---------------------------------------------------------------|------------------------------------------------------------------------------------------------|--------------------------------------------------------------------------------------------------------------------------------------------------------------------------|-------------------------------------------------------------------------------------------|-----------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| | | Long Term (LT) vs. Short Term (ST) | Fixed Rent (FR) vs. Sharecropping (SC) | | |
| of the ction | Current output depends on the tenant's non observable effort. | | FR provides stronger incentives for production effort. | Crop's sensitivity to effort. | Other things equal, FR is more likely to be chosen the higher the crop's sensitivity to effort and the lower the tenant's cost of effort. |
| Characteristics of the Production Function | | LT provides stronger incentives for investment effort. | Under multi-tasking, FR provides strong incentives for production effort at the expense of investment effort. | Crop's sensitivity to investment effort. | Other things equal, LT is more likely the higher the crop's sensitivity to investment and the lower the tenant's cost of investment. Because of multitasking, FR is less likely for crops that are more sensitive to investment. |
| | liability and credit markets are | By committing to LT, the landlord gives up eviction threats as an instrument to elicit effort. | Limited liability imposes an upper bound to the feasible rent payment in a FR contract. SC allows the landlord to extract more surplus through statecontingent payments. | returns between the good and the bad | Other things equal, FR/LT is more likely to be chosen for rich tenant, for crops with low return spread and when the outside option of the tenant is high. |
| Characteristics of the Tenant and Credit Markets | markets are imperfect. | LT guarantees employment and allows the tenant to smooth consumption. | | Tenant's wealth and crop's riskiness determines the trade-off between risk and insurance. | Other things equal, LT is more likely to be chosen for poor tenants and risky crops. FR is more likely to be chosen for rich tenants and low-risk crops. |
| Transaction Costs | | negotiated every year. On the | FR entails lower transaction costs because under SC the landlord must monitor output measurement and division to prevent opportunism. | Landlord's opportunity cost of time and flexibility needs. | Other things equal, LT is more likely to be chosen when the landlord's opportunity cost of time is high and the need for flexibility is low. FR is more likely to be chosen when the landlord's opportunity cost of time is high. |

TABLE 2 DESCRIPTIVE STATISTICS

PART A: CONTRACTS

| | fixed rent | share | total |
|------------|------------|-------|-------|
| long term | 435 | 85 | 520 |
| long term | (62) | (12) | (74) |
| short term | 166 | 19 | 185 |
| SHORE CHIL | (24) | (3) | (26) |
| total | 601 | 104 | 705 |
| lotai | (85) | (15) | (100) |
| | | | |

Source: State and Notary Archives, Siracusa, Italy. Volumes and years reported in table A1. Numbers in parenthesis are percentages of the total sample.

PART C: TENANTS AND CONTRACTS

| | poor | middle class | rich |
|-----------------------|------|-----------------|------|
| long term/fixed rent | 100 | 176 | 159 |
| long terminixed rent | (42) | (62) | (88) |
| long term/ share | 39 | 34 | 12 |
| long termi share | (16) | (12) | (7) |
| short term/fixed rent | 91 | 68 | 7 |
| Short terminixed rent | (38) | (24) | (4) |
| short term/share | 8 | 8 | 3 |
| Short termishare | (3) | (3) | (2) |
| total | 238 | 286 | 181 |

Number in parenthesis are percentages of the total by class. For instance, 42% of contracts with poor tenants are long term/ fixed rent.

PART B: CROPS AND CONTRACTS

| | annual | all trees | vines & citrus trees |
|-----------------------|--------|-----------|-------------------------|
| long term/fixed rent | 94 | 341 | 194 |
| long term/lixed rent | (36) | (76) | (78) |
| long term/ share | 15 | 70 | 50 |
| long term/ share | (6) | (16) | (20) |
| short term/fixed rent | 137 | 29 | 4 |
| Short term/lixed rent | (53) | (6.5) | (2) |
| short term/share | 13 | 6 | 2 |
| SHOIL LEITH/SHATE | (5) | (1.5) | (1) |
| total | 259 | 196 | 250 |

Number in parenthesis are percentages of the total by crop. For instance, 36% of annual crop contracts are long term/ fixed rent.

PART D: LANDLORDS AND CONTRACTS

| | absentee | aristocrat | female | all landlords |
|-----------------------|----------|------------|--------|------------------|
| long term/fixed rent | 77 | 28 | 110 | 435 |
| | (43) | (88) | (86) | (62) |
| long term/ share | 14 | 2 | 8 | 85 |
| | (8) | (6) | (6) | (12) |
| short term/fixed rent | 89 | 2 | 8 | 166 |
| | (49) | (6) | (6) | (24) |
| short term/share | 0 | 0 | 2 | 19 |
| | (0) | (0) | (2) | (3) |
| total | 180 | 32 | 128 | 705 |

Number in parenthesis are percentages of the total by class. For instance, 43% of contracts signed by landlords who live in a different town are long term/ fixed rent.

TABLE 3. DETERMINANTS OF CONTRACTUAL STRUCTURE

Bivariate Probit Estimates

Dependent Variables: contract duration (cols. a) and contract type (cols. b)

Standard Errors in parenthesis, marginal effects (evaluated at sample mean) in brackets.

| | duration | type | duration | type | duration | type | duration | type |
|-------------------------------------|----------|--------|----------|--------|----------|-----------|----------|---------|
| | (1a) | (1b) | (2a) | (2b) | (3a) | (3b) | (4a) | (4b) |
| crop is tree | 1.61*** | 286** | 1.50*** | 379*** | 1.39*** | 296* | .975*** | 044 |
| | (.117) | (.125) | (.119) | (.136) | (.132) | (.154) | (.210) | (.211) |
| | [.500] | [063] | [.452] | [079] | [.405] | [060] | [.144] | [004] |
| tenant is poor | | | -1.03*** | 621*** | -1.02** | 646*** | 547*** | 780*** |
| | | | (.210) | (.176) | (.219) | (.187) | (.228) | (.207) |
| | | | [314] | [153] | [298] | [153] | [074] | [085] |
| tenant is middle class | | | 728*** | 375** | 813*** | 380** | 160 | 592*** |
| | | | (.211) | (.169) | (.226) | (.170) | (.237) | (.187) |
| | | | [208] | [087] | [223] | [084] | [019] | [056] |
| landlord is female | | | | | .888*** | .502*** | .635*** | .561*** |
| | | | | | (.213) | (.178) | (.207) | (.208) |
| | | | | | [.172] | [.088] | [.055] | [.035] |
| landlord belongs to the aristocracy | | | | | 1.04** | .476 | 1.36*** | .659* |
| | | | | | (.527) | (.357) | (.471) | (.410) |
| | | | | | [.161] | [.078] | [.064] | [.033] |
| landlord lives in a different town | | | | | 360*** | .458*** | 234 | .310 |
| | | | | | (.137) | (.182) | (.236) | (.213) |
| | | | | | [100] | [.085] | [029] | [.023] |
| correlation coefficient | 0 | 82 | 0 | 98 | 0 | 95 | .0 | 79 |
| wald test, p-value | .38 | 86 | .3 | 15 | .34 | 43 | .5 | 20 |
| town controls | n | 0 | n | 0 | n | 0 | y | es |
| joint F-test, p-value | | | | | | | .000 | .000 |
| year controls | n | 0 | no | | no | | y | es |
| joint F-test, p-value | | | | | | .000 .000 | | .000 |
| joint log likelihood | -59 | 0.8 | -56 | 5.2 | -53 | 7.2 | -39 | 6.17 |

Notes. Source: see T2. **Number of observations** is 705 in all specifications. **Standard Errors** are based on White (1982)'s robust "sandwich" estimator for the asymptotic covariance matrix. **Marginal Effects** are computed as the change in the probability of the positive outcome when the variable of interest changes from 0 to 1. The **Wald test** statistics is the t ratio of the correlation coefficient squared and has a chi-squared distribution with 1 degree of freedom. **Omitted categories** are: *annual* for crop type and *rich* for tenant's class.*, **, and *** indicates significance at the 10%,5% and 1% level.

TABLE 4. MARGINAL EFFECTS ON THE JOINT PROBABILITIES

Bivariate probit estimates from cols. 4a and 4b, Table 4

Marginal effects evaluated at sample mean. Standard errors in parenthesis.

| | | Joint Probabilities | Marginal P | robabilities | | |
|-------------------------------------|---------------------------|-----------------------|----------------------------|-----------------------|---------------|----------------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| | Pr(long term, fixed rent) | Pr (long term, share) | Pr(short term, fixed rent) | Pr(short term, share) | Pr(long term) | Pr(fixed rent) |
| crop is tree | .134* | .010 | 137** | 006** | .144*** | 004 |
| | (.069) | (.012) | (.068) | (.003) | (.041) | (.017) |
| tenant is poor | 145*** | .071* | .060 | .014** | 074** | 084*** |
| | (.052) | (.036) | (.040) | (.006) | (.035) | (.029) |
| tenant is middle class | 070* | .051* | .013 | .006** | 019 | 056*** |
| | (.036) | (.026) | (.023) | (.003) | (.028) | (.022) |
| landlord is female | .085*** | 031** | 050* | 004** | .055*** | .035*** |
| | (.029) | (.012) | (.027) | (.002) | (.015) | (.009) |
| landlord belongs to the aristocracy | .093*** | 029** | 060* | 004** | .064*** | .033*** |
| | (.034) | (.014) | (.031) | (.002) | (.012) | (.011) |
| landlord lives in a different town | 007 | 022* | .031 | 001 | 029 | .023 |
| | (.030) | (.013) | (.028) | (.001) | (.031) | (.014) |
| town controls | yes | yes | yes | yes | yes | yes |
| year controls | yes | yes | yes | yes | yes | yes |

Source: see T.2. **Marginal effects** are computed from the bivariate probit estimates in cols 4a and 4b, Table 3. **Standard Errors** are bootstrapped using 1000 replications. **Omitted categories** are: annual for crop type and rich for tenant's class. *, **, and *** indicate significance at the 10%,5% and 1% level.

TABLE 5. DETERMINANTS OF CONTRACTUAL STRUCTURE: INTERACTIONS

Bivariate Probit Estimates

Dependent Variables: contract duration (col 1a) and contract type (cols 1b)
Standard Errors in parenthesis, marginal effects (evaluated at sample mean) in brackets.

| | (1a) | (1b) |
|-------------------------------------|---------|---------|
| crop is tree | .472 | .528 |
| | (.495) | (.371) |
| | [860.] | [.052] |
| tenant is poor | -1.09** | 224 |
| | (.477) | (.370) |
| | [187] | [020] |
| tenant is middle class | 214 | 001 |
| | (.478) | (.418) |
| | [029] | [001] |
| crop is tree*tenant is poor | .875* | 730* |
| | (.524) | (.429) |
| | [.075] | [095] |
| crop is tree*tenant is middle class | .043 | 764* |
| | (.534) | (.473) |
| | [.006] | [092] |
| landlord is female | .671*** | .539*** |
| | (.204) | (.209) |
| | [.064] | [.034] |
| landlord belongs to the aristocracy | 1.34*** | .681* |
| | (.467) | (.405) |
| | [.072] | [.034] |
| landlord lives in a different town | 161 | .313 |
| | (.231) | (.213) |
| | [022] | [.023] |
| test 1, p-value | .415 | |
| correlation coefficient | .05 | 54 |
| wald test, p-value | .65 | 54 |
| town controls | ye | |
| year controls | ye | |
| joint log likelihood | -39 | 1.7 |

Source: see T2. **Notes:** see T4. The null hypothesis for Test 1 is that the coefficients of "tree" and that of "tree*poor" are equal.

TABLE 6 JOINT ESTIMATES OF LENGTH (Continuous) AND TYPE

FIML estimates: joint ordered probit (for length) and probit (for type) estimates Dependent Variables: contract length (cols. 1a,2a) and contract type (cols. 1b,2b) Standard Errors in parenthesis.

| | duration | type | duration | type | |
|-------------------------------------|----------|---------|----------|---------|--|
| | (1a) | (1b) | (2a) | (2b) | |
| crop is tree | .678*** | 042 | .465* | .522 | |
| | (.143) | (.211) | (.242) | (.371) | |
| tenant is poor | 238* | 782*** | 646** | 237 | |
| | (.130) | (.206) | (.280) | (.372) | |
| tenant is middle class | 018 | 591*** | .053 | 002 | |
| | (.105) | (.186) | (.277) | (.418) | |
| crop is tree*tenant is poor | | | .597** | 714* | |
| | | | (.303) | (.431) | |
| crop is tree*tenant is middle class | | | 094 | 761* | |
| | | | (.293) | (.473) | |
| landlord is female | 123 | .560*** | 116 | .538*** | |
| | (.106) | (.208) | (.105) | (.209) | |
| landlord belongs to the aristocracy | .687*** | .647 | .665*** | .668* | |
| | (.222) | (.407) | (.232) | (.402) | |
| landlord lives in a different town | 052 | .299 | 024 | .302 | |
| | (.141) | (.213) | (.138) | (.213) | |
| correlation coefficient | .0 | 59 | .0: | 57 | |
| wald test, p-value | .4 | 21 | .43 | 32 | |
| town controls | y | es | y€ | es | |
| year controls | y | es | yes | | |
| joint log likelihood | -124 | 8.27 | -124 | 1.55 | |

Source: see T2. Notes: see T4.

TABLE 7. MARGINAL EFFECTS ON DURATION OUTCOMES

7.1 BASIC SPECIFICATION (from column 1a, Table 6)

| | Pr(d=1) | Pr(d=2) | Pr(d=3) | Pr(d=4) | Pr(d=5) | Pr(d=6) | Pr(d=7) | Pr(d=8) | Pr(d=9) | Pr(d=10) |
|-------------------------------------|-----------|-----------|-----------|----------|----------|----------|----------|----------|----------|----------|
| crop is tree | -0.209*** | -0.040*** | -0.016*** | 0.141*** | 0.055*** | 0.049*** | 0.004** | 0.008*** | 0.006** | 0.0017 |
| | (0.046) | (0.009) | (0.006) | (0.032) | (0.012) | (0.010) | (0.002) | (0.003) | (0.002) | (0.0012) |
| tenant is poor | 0.071* | 0.016* | 0.008* | -0.049* | -0.020* | -0.018* | -0.001 | -0.003* | -0.002 | -0.0006 |
| | (0.040) | (0.009) | (0.004) | (0.028) | (0.011) | (0.009) | (0.001) | (0.0017) | (0.0015) | (0.0005) |
| tenant is middle class | 0.005 | 0.001 | 0.0007 | -0.004 | -0.002 | -0.001 | -0.0001 | -0.0002 | -0.0002 | -0.0001 |
| | (0.031) | (0.007) | (0.004) | (0.021) | (0.009) | (800.0) | (0.0007) | (0.001) | (0.001) | (0.0003) |
| landlord is female | 0.037 | 0.008 | 0.004 | -0.026 | -0.011 | -0.009 | -0.0008 | -0.0016 | -0.001 | -0.0003 |
| | (0.033) | (0.007) | (0.003) | (0.023) | (0.009) | (800.0) | (0.0007) | (0.0013) | (0.001) | (0.0003) |
| landlord belongs to the aristocracy | -0.149*** | -0.054*** | -0.051** | 0.065*** | 0.065*** | 0.077** | 0.008 | 0.017* | 0.016 | 0.006 |
| | (0.034) | (0.018) | (0.022) | (0.014) | (0.022) | (0.033) | (0.005) | (0.010) | (0.010) | (0.005) |
| landlord lives in a different town | 0.015 | 0.004 | 0.002 | -0.010 | -0.004 | -0.004 | -0.0003 | -0.0007 | -0.0005 | -0.0001 |
| | (0.041) | (0.010) | (0.005) | (0.028) | (0.012) | (0.011) | (0.0009) | (0.0018) | (0.001) | (0.0004) |

7.2 INTERACTIONS (from column 2a, Table 6)

| | Pr(d=1) | Pr(d=2) | Pr(d=3) | Pr(d=4) | Pr(d=5) | Pr(d=6) | Pr(d=7) | Pr(d=8) | Pr(d=9) | Pr(d=10) |
|-------------------------------------|---------|---------|---------|---------|---------|---------|---------|---------|---------|----------|
| crop is tree | 141* | 030** | 012** | .097* | .038** | .034** | .003 | .006* | .004 | .001 |
| | (.076) | (.014) | (.005) | (.052) | (.019) | (.017) | (.002) | (.003) | (.003) | (.009) |
| tenant is poor | .201** | .038*** | .013** | 138** | 051** | 044** | 0036* | 007** | 005* | 001 |
| | (.092) | (.013) | (.006) | (.062) | (.021) | (.018) | (.0021) | (.003) | (.003) | (.001) |
| tenant is middle class | 015 | 004 | 002 | .010 | .004 | .004 | .0003 | .0007 | .0005 | .0001 |
| | (.080.) | (.019) | (.010) | (.054) | (.024) | (.024) | (.0018) | (.004) | (.0023) | (.0007) |
| crop is tree*tenant is poor | 146*** | 046** | 037 | .082*** | .055 | .059 | .005 | .012 | .010 | .003 |
| | (.060) | (.025) | (.025) | (.022) | (.030) | (.038) | (.004) | (.009) | (800.) | (.003) |
| crop is tree*tenant is middle class | .028 | .006 | .003 | 019 | 008 | 007 | 0006 | 001 | 001 | 0002 |
| | (.088) | (.019) | (.009) | (.061) | (.024) | (.021) | (.002) | (.003) | (.002) | (.0006) |
| landlord is female | .034 | .008 | .003 | 024 | 009 | 009 | 0007 | 001 | 001 | 0002 |
| | (.032) | (.007) | (.003) | (.022) | (800.) | (800.) | (.0007) | (.001) | (.001) | (.0003) |
| landlord belongs to the aristocracy | 141*** | 053*** | 047** | .068*** | .063 | .073** | .007 | .016* | .014 | .005 |
| | (.036) | (.019) | (.023) | (.011) | (.022) | (.034) | (.005) | (.009) | (.009) | (.004) |
| landlord lives in a different town | .007 | .001 | .0008 | 005 | 073 | 001 | 0001 | 0003 | 0002 | 0001 |
| | (.040) | (.009) | (.005) | (.027) | (.011) | (.010) | (8000.) | (.001) | (.001) | (.0003) |

Source: see T.2. Marginal effects are computed from the ordered probit estimates in cols 1a and 2a Table 6.Standard Errors are bootstrapped using 1000 replications.Omitted categories are: annual for crop type and rich for tenant's class. *, **, and *** indicate significance at the 10%,5% and 1% level.

TABLE 8. EXTENSIONS AND ALTERNATIVE VARIABLES DEFINITIONS

Bivariate Probit Estimates, Dependent Variables: contract duration (cols. a) and contract type (cols. b) Standard Errors in parenthesis, marginal effects (evaluated at sample mean) in brackets.

| | duration | type | duration | type | duration | type |
|------------------------------------------|----------|---------|----------|---------|----------|---------|
| | (1a) | (1b) | (2a) | (2b) | (3a) | (3b) |
| crop is tree | | | | | 1.17*** | 117 |
| | | | | | (.200) | (.199) |
| | | | | | [.261] | [018] |
| crop is olive or fruit tree | .570*** | .301 | | | | |
| | (.224) | (.234) | | | | |
| | [.046] | [.021] | | | | |
| crop is vines or citrus tree | 1.61*** | 319 | | | | |
| | (.266) | (.232) | | | | |
| | [.131] | [028] | | | | |
| crop is tree (either main or secondary) | | | 1.56*** | .271 | | |
| | | | (.214) | (.210) | | |
| | | | [.265] | [.026] | | |
| tenant is poor | 537*** | 817*** | 466** | 763*** | 643*** | 783*** |
| | (.231) | (.205) | (.226) | (.206) | (.218) | (.197) |
| | [062] | [086] | [054] | [085] | [136] | [147] |
| tenant is middle class | 213 | 562*** | 055 | 596*** | 280 | 618*** |
| | (.247) | (.189) | (.243) | (.187) | (.231) | (.174) |
| | [022] | [051] | [005] | [058] | [054] | [106] |
| landlord is female | .618*** | .583*** | .719*** | .574*** | .511** | .651*** |
| | (.219) | (.214) | (.476) | (.211) | (.208) | (.206) |
| | [.044] | [.034] | [.051] | [.037] | [.077] | [.077] |
| landlord belongs to the aristocracy | 1.57*** | .711* | 1.42*** | .702* | .968** | .608 |
| | (.473) | (.407) | (.476) | (.424) | (.489) | (.384) |
| | [.054] | [.032] | [.055] | [.035] | [.010] | [.064] |
| landlord lives in a different town | 279 | .335 | 214 | .346 | 439** | .323 |
| | (.243) | (.208) | (.239) | (.214) | (.187) | (.209) |
| | [031] | [.023] | [024] | [.026] | [093] | [.045] |
| daily wage for adult male casual workers | | | | | 601 | 1.06** |
| | | | | | (.636) | (.511) |
| | | | | | [113] | [.166] |
| test 1, p-value | .000 | .000 | | | | |
| correlation coefficient | .221 | .221 | 012 | 012 | .032 | .032 |
| wald test, p-value | .075 | .075 | .924 | .924 | .774 | .774 |
| town controls | yes | yes | yes | yes | no | no |
| region controls | no | no | no | no | yes | yes |
| year controls | yes | yes | yes | yes | yes | yes |
| joint log likelihood | -377.6 | -377.6 | -381.7 | -381.7 | -433.2 | -433.2 |

Source: see T2. The null hypothesis for Test 1 is that the coefficients of "olives and fruit" and that of "vines and citrus" are equal.

TABLE 9. MATCHING OF TENANTS AND CROPS

Probit Estimates; Standard Errors in parenthesis.

| | crop is tree | crop is vines or citrus |
|------------------------|--------------|-------------------------|
| tenant is poor | 228 | 076 |
| | (.193) | (.160) |
| tenant is middle class | .159 | .194 |
| | (.199) | (.144) |
| town controls | yes | yes |
| year controls | yes | yes |
| log likelihood | -211.5 | -353.1 |

Source: see T2. Standard Errors are based on White (1982)'s robust "sandwich" estimator for the asymptotic covariance matrix.